# Lesson 16: Forecasting Stationary Time Series

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- to understand the structure of the time series
- to predict future values of the time series

In this lesson, we consider the second goal:

### to predict future values of a time series

Let  $\{x_t\}$  be a stationary process. Let h > 0 be an integer.

We consider the problem of predicting the value  $x_{T+h}$  in terms of the values

$$\Omega_{\mathcal{T}} = \{x_{\mathcal{T}}, x_{\mathcal{T}-1}...\}$$

up time T.

Let us denote this *h*-step-ahead forecast at time T by  $\hat{x}_{T,h}$ .

Any function of the random variables  $x_T, x_{T-1}...$ , can be considered like an *h*-step-ahead prediction.

#### Which is the best?

The accuracy of forecasts is evaluated by some loss function,  $L(x_{T+h}, \hat{x}_{T,h})$ , that represents the penality or cost when we predict  $\hat{x}_{T,h}$  but the outcome is  $x_{T+h}$ .

A typical loss function is the **Mean Square Error** (**MSE**) defined by

$$MSE(\hat{x}_{T,h}) = E[(x_{t+h} - \hat{x}_{T,h})^2]$$

We suppose that the forecaster's objective is to minimize the MSE.

It can be shown that the optimal (minimum MSE) *h*-step-ahead forecast of  $x_{T+h}$  at time T is the **conditional expectation** 

 $E\left[x_{T+h}|\Omega_{T}\right]$ 

In other terms, if  $\hat{x}_{\mathcal{T},h}$  is any h-step predictor at time  $\mathcal{T},$  we have that

 $\mathsf{MSE}(\hat{x}_{T,h}) \geq \mathsf{MSE}(E[x_{T+h}|\Omega_T])$ 

Consider a causal ARMA(p, q) process

$$x_{t} = \nu + \phi_{1}x_{t-1} + \dots + \phi_{p}x_{t-p} + u_{t} + \theta_{1}u_{t-1} + \dots + \theta_{q}u_{t-q}$$

To simplify the discussion, it is assumed that the actual coefficients

$$\{\nu,\phi_1,...,\phi_p,\theta_1,...,\theta_q\}$$

and current and past realizations of  $x_t$  and  $u_t$  are known to the researcher.

#### The process is causal so that it has an MA representation

$$x_t = \mu + u_t + \psi_1 u_{t-1} + \dots = \mu + \sum_{i=0}^{\infty} \psi_i u_{t-i}$$

with

$$\mu = \frac{\nu}{1 - \sum_{i=0}^{p} \phi_i}.$$

We also assume that our process is invertible so that it has an AR representation

$$x_t = \gamma + \pi_1 x_{t-1} + \pi_2 x_{t-2} + \dots + u_t = \gamma + \sum_{i=1}^{\infty} \pi_i x_{t-i} + u_t$$

with

$$\gamma = \left(1 + \sum_{i=1}^{\infty} \pi_i\right) \mu$$

The AR and MA representations show that for this process the information in

$$\{x_T, x_{T-1}...\}$$

can be equivalently be represented as

$$\{u_T, u_{T-1}...\}$$

since each  $u_t$  can be computated from the past and present  $x_s$ ,  $s \le t$  and each  $x_t$  can be obtained from the past and present  $u_s$ ,  $s \le t$ .

Assuming in addition that  $u_t$  and  $u_s$  are independent and not only uncorrelated for  $s \neq t$  we can now determine the conditional expectation  $E[x_{T+h}|\Omega_T]$ , that is optimal (minimum MSE) *h*-step-ahead forecast of  $x_{T+h}$  at time *T*. For istance, for h = 1

$$\begin{aligned} \hat{x}_{T,1} &= E \left[ x_{T+1} | \Omega_T \right] = \nu + \phi_1 E \left[ x_T | \Omega_T \right] + \dots + \phi_p E \left[ x_{T+1-p} | \Omega_T \right] \\ &+ E \left[ u_{T+1} | \Omega_T \right] + \theta_1 E \left[ u_T | \Omega_T \right] + \dots + \theta_q E \left[ u_{T+1-q} | \Omega_T \right] \\ &= \nu + \phi_1 x_T + \dots + \phi_p x_{T+1-p} + \theta_1 u_T + \dots + \theta_q u_{T+1-q} \end{aligned}$$

#### For an arbitrary positive h we get

$$\hat{x}_{T,h} = \nu + \phi_1 \hat{x}_{T,h-1} + \ldots + \phi_p \hat{x}_{T,h-p} + \theta_h u_T + \ldots + \theta_q u_{T+h-q} \text{ if } h \leq q$$

and

$$\hat{x}_{T,h} = \nu + \phi_1 \hat{x}_{T,h-1} + \ldots + \phi_p \hat{x}_{T,h-p}$$
 if  $h > q$ 

With these formulas, forecast can be computed recursively.

To illustrate these formulas we consider the MA(3) process

$$x_{t} = \nu + u_{t} + \theta_{1}u_{t-1} + \theta_{2}u_{t-2} + \theta_{3}u_{t-3}$$

For this process

$$\hat{x}_{T,1} = \nu + \theta_1 u_T + \theta_2 u_{T-1} + \theta_3 u_{T-2}$$

$$\hat{x}_{T,2} = \nu + \theta_2 u_T + \theta_3 u_{T-1}$$

$$\hat{x}_{T,3} = \nu + \theta_3 u_T$$

$$\hat{x}_{T,h} = \nu \quad \forall h \ge 4$$

As the MA(3) process has a memory of only 3 periods, all forecasts 4 or more steps ahead collapse to the intercept (the mean).

### Now, we consider a causal AR(2) process

$$x_t = \nu + \phi_1 x_{t-1} + \phi_2 x_{t-2} + u_t$$

#### For this process

$$\hat{x}_{T,1} = \nu + \phi_1 x_T + \phi_2 x_{T-1}$$
$$\hat{x}_{T,2} = \nu + \phi_1 \hat{x}_{T,1} + \phi_2 x_T$$
$$\hat{x}_{T,3} = \nu + \phi_1 \hat{x}_{T,2} + \phi_2 \hat{x}_{T,1}$$

$$\hat{x}_{T,h} = \nu + \phi_1 \hat{x}_{T,h-1} + \phi_2 \hat{x}_{T,h-2} \quad \forall h \ge 3$$

Alternatively, the optimal predictor can be determined using the AR or MA representation of  $x_t$ 

$$\hat{x}_{\mathcal{T},h} = \gamma + \sum_{i=1}^{\infty} \pi_i \hat{x}_{\mathcal{T},h-i}$$

$$\hat{x}_{T,h} = \mu + \sum_{i=h}^{\infty} \psi_i u_{T+h-i}$$

From the formula

$$\hat{x}_{T,h} = \mu + \sum_{i=h}^{\infty} \psi_i u_{T+h-i}$$

the forecast error is easy to obtain.

We have

$$e_{x,T+h} = x_{T+h} - \hat{x}_{T,h} = \sum_{i=0}^{h-1} \psi_i u_{T+h-i}$$

The forecast is unbiased since the expected error is zero

$$E[x_{T+h} - \hat{x}_{T,h}] = \sum_{i=0}^{h-1} \psi_i E[u_{T+h-i}] = 0$$

The variance of the forecast error is

$$\sigma_h^2 = E\left[\left(\sum_{i=0}^{h-1} \psi_i u_{T+h-i}\right)^2\right] = \sigma_u^2 \sum_{i=0}^{h-1} \psi_i^2$$

We note that if  $u_t \sim i.i.N(0,\sigma^2)$ , then

$$rac{x_{\mathcal{T}+h}-\hat{x}_{\mathcal{T},h}}{\sigma_h}\sim {\sf N}(0,1)$$

where  $\sigma_h$  is the square root of  $\sigma_h^2$ .

Denoting by  $z_{\alpha}$  the upper  $\alpha$ 100 percentage point of the standard normal distribution we get

$$1 - \alpha = P\left(-z_{\alpha/2} \leq \frac{x_{T+h} - \hat{x}_{T,h}}{\sigma_h} \leq z_{\alpha/2}\right).$$

Hence, a  $(1 - \alpha)100\%$  interval forecast *h* periods ahead for  $x_{T+h}$  is

$$\left[\hat{x}_{T,h}-z_{\alpha/2}\sigma_h,\hat{x}_{T,h}+z_{\alpha/2}\sigma_h\right].$$

The meaning of this interval is the following. If the forecast interval is computed repeatedly from a large number of realizations of the considered stochastic process, then  $(1 - \alpha)100\%$  of the intervals will contain the actual value (the realization) of the random variable  $x_{T+h}$ .

Following Granger and Newbold (1976), we define as a measure of predictability of a stochastic process  $x_t$ ;  $t \in \mathbb{Z}$ , with finite variance, the index

$$R_x^2 = 1 - \frac{\mathsf{var}(e_{x,T+1})}{\mathsf{Var}(x_t)}$$

We note that  $R_x^2 \in [0, 1]$ .

- If  $R_x^2 = 0$  the process  $x_t \ t \in \mathbb{Z}$  is unpredictable;
- If  $R_x^2 = 1$  the process  $x_t \ t \in \mathbb{Z}$  is perfectly predictable.

Now, we suppose that  $x_t \ t \in \mathbb{Z}$  is an ARMA(p, q) causal and invertible stochastic process,

$$x_{t} = \nu + \phi_{1}x_{t-1} + \dots + \phi_{p}x_{t-p} + u_{t} + \theta_{1}u_{t-1} + \dots + \theta_{q}u_{t-q}, \quad u_{t} \sim WN(0, \sigma^{2})$$

The innovation  $u_{T+1}$  is the corresponding one-step forecast error. Thus we have

$$R_x^2 = 1 - rac{\sigma^2}{\mathsf{Var}(x_t)}$$

The index  $R_x^2$  can be related to the coefficients  $\{\psi_1, \psi_2, ...\}$  of the MA representation of  $x_t$ .

## Predictability of a stochastic process

Since

$$\operatorname{Var}(x_t) = \left(1 + \sum_{j=1}^{\infty} \psi_j^2\right) \sigma^2,$$

we obtain

$$R_x^2 = 1 - \frac{1}{1 + \sum_{j=1}^{\infty} \psi_j^2} = \frac{\sum_{j=1}^{\infty} \psi_j^2}{1 + \sum_{j=1}^{\infty} \psi_j^2}.$$

We call the sequence  $\{\psi_j\} \psi$ -weights; they represent the dependence structure of the series.

A series with small  $\psi$ -weights (with a few structure) will be less predictable than one with large  $\psi$ -weights (a series with more structure). Thus, this predictability measure provides a synthetic evaluation of the dependence structure of a stationary time series. In general, we can define a measure of predictability relative to a h-steps forecast by

$$R_x^2(h) = 1 - \frac{\mathsf{Var}(e_{x,T+h})}{\mathsf{Var}(x_t)} = 1 - \frac{1 + \sum_{j=1}^{h-1} \psi_j^2}{1 + \sum_{j=1}^{\infty} \psi_j^2} = \frac{\sum_{j=h}^{\infty} \psi_j^2}{1 + \sum_{j=1}^{\infty} \psi_j^2}.$$

This predictability index is utilised by Hong and Billings (1999) and for h = 1 coincides with the index proposed in Granger and Newbold (1976).

Suppose that

$$\hat{x}_{T+h} = g(x_T, x_{T-1}, ...; \hat{\theta}),$$

is a forecasting methods used to forecast  $x_{T+h}$  for h = 1, 2, ..., N. How good is the forecast?

After we have observed the real values  $x_{T+h}$ , h = 1, 2, ..., N, we can calculate the forecast errors

$$e_{x,T+h} = x_{T+h} - \hat{x}_{T+h}$$
  $h = 1, 2, ..., N$ 

Various accuracy measures, based on these errors, have been used to evaluate the performance of forecasting methods. We will present three performance measures

- mean squared error (MSE)
- root mean squared error (RMSE)
- mean absolute percent error (MAPE)

The most common measure is the mean squared error (MSE). It is defined as

$$MSE = \frac{1}{N} \sum_{h=1}^{N} e_{x,T+h}^2$$

A widely used measure of overall accuracy of a forecasting method is the root mean squared error (RMSE). The RMSE statistic is defined as follows

$$RMSE = \sqrt{\frac{\sum_{h=1}^{N} e_{x,T+h}^2}{N}}$$

Another common criterion used in comparing performance of forecast models is the mean absolute percent error (MAPE).The MAPE statistic is defined as follows

$$MAPE = \frac{100}{N} \sum_{h=1}^{N} \left| \frac{e_{x,T+h}}{x_{T+h}} \right|$$

The MSE, RMSE, and MAPE measure the magnitude of the forecast errors. Better models will show smaller values for these statistics.

Each of these measurements has different advantages and limitations. Often the square root of MSE (RMSE) is used so as to preserve the units. RMSE, however, does not provide information about the relative magnitude of the forecast error. Hence, using more than one performance measure is always recommended.

Some example of forecasts:

- "I think there is a world market for about five computers" -Founder of IBM in 1947
- "There is no reason for any individual to have a computer in their home" President Digital Equipment in 1977
- "Stock prices have reached what looks like a permanently high plateau" Yale Professor of Economics in September 1929

#### It's not easy to make good forecasts!

However, these are judgmental forecasts. In this lecture we are interested to the ARMA time series forecasts.

#### Are the ARMA time series forecasts reliable?

A reliable ARMA time series forecast requires that the future is not too different from the past.

In other terms, it requires that  $x_{T+h}$  will be drawn from the same DGP that generated the previous observations.

## The reliability of ARMA time series forecasts



If the future is too different from the past the ARMA model will produce biased forecasts.

In particular, if a structural break will occur between T and T + h, then  $x_{T+h}$  will be drawn from a different DGP from that has generated the previous observations.

## The reliability of ARMA time series forecasts



### The ARMA model is mis-specified for new DGP and then any forecast, based on this model, is wrong.

#### Advantages:

- constitutes a flexible class of models;
- 2 provide unconditional forecasts;
- Image of the second second

#### Disadvantages:

- requires large number of observations for model identification (at least 50 and preferably 100 observations should be available to build a proper model);
- 2 need a long series of data without structural change;
- the amount of subjective input at the identification stage make them somewhat more of an art than a science